

Mediation Analysis: OLS vs. SUR vs. ISUR vs. 3SLS vs. SEM

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In Chap. 11 of “Statistical Analysis of Management Data” (Gatignon, 2014), tests of mediation are presented using the SUR estimation method with estimation of indirect effects through the bootstrap procedure. In this additional note, the appropriateness of SUR versus OLS or 3SLS is discussed.

Case of single exogenous (iv) variable (with single or multiple mediators)

The model is:

$$iv \rightarrow med \rightarrow dv$$

In equation:

$$med_i = \alpha_0 + \alpha_1 iv_i + \varepsilon_{1i} \quad (1)$$

$$dv_i = \beta_0 + \beta_1 med_i + \varepsilon_{2i} \quad (2)$$

If $\rho(\varepsilon_{1i}, \varepsilon_{2i}) = 0$, then OLS and SUR provide the same estimates. Note that in practice, the estimate of the correlation will not be exactly zero and the estimates may vary slightly. This is possible because 1) the generalized least squares estimator is the same as the OLS estimator and 2) the independent variable in Eq. (2) is not correlated to the error term of the equation, which does not lead to a biased estimator.

The same reasoning applies to the extended second equation where the direct effect of iv is included:

$$dv_i = \beta_0 + \beta_1 med_i + \beta_2 iv_i + \varepsilon_{2i} \quad (3)$$

To illustrate this, we created a dataset generated by the following data generating functions:

$$\begin{aligned} med_i &= 12 - 0.6iv_i + \varepsilon_{1i} \quad ; \quad \sigma_{\varepsilon_1} = 1.5 \\ dv_i &= 5 + 0.75med_i + 0.1iv_i + \varepsilon_{2i} \quad ; \quad \sigma_{\varepsilon_2} = 3 \end{aligned} \quad (4)$$

The OLS estimates are provided in Table 1.

Table 1 – OLS estimates

```
. use "/Users/gatignon/Documents/WORK_STATA/MediationData/mediationdata_singleiv_Rho_0.dta",
clear

. regress med iv
```

Source	SS	df	MS	Number of obs = 400		
Model	582.630641	1	582.630641	F(1, 398)	=	263.44
Residual	880.233943	398	2.21164307	Prob > F	=	0.0000
				R-squared	=	0.3983
				Adj R-squared	=	0.3968
Total	1462.86458	399	3.66632728	Root MSE	=	1.4872

med	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
iv	-.6360555	.0391883	-16.23	0.000	-.7130973	-.5590137
_cons	12.30566	.3964391	31.04	0.000	11.52629	13.08504


```
. regress dv med iv
```

Source	SS	df	MS	Number of obs = 400		
Model	947.534487	2	473.767243	F(2, 397)	=	56.21
Residual	3346.36766	397	8.42913767	Prob > F	=	0.0000
				R-squared	=	0.2207
				Adj R-squared	=	0.2167
Total	4293.90214	399	10.7616595	Root MSE	=	2.9033

dv	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
med	.8976328	.0978571	9.17	0.000	.7052498	1.090016
iv	.1641715	.0986263	1.66	0.097	-.0297237	.3580667
_cons	3.390934	1.431462	2.37	0.018	.5767419	6.205127

As can be seen, all parameter estimates are close to the true parameters. The SUR estimates are not needed in this case where the error terms across equations are uncorrelated. Because, a priori, this information is not known, it can be seen in Table 2 that the SUR parameter estimates are almost identical to the OLS estimates.

Table 2 – SUR Estimates

```
. sureg (med iv) (dv med iv), corr
```

Seemingly unrelated regression

Equation	Obs	Parms	RMSE	"R-sq"	chi2	P
med	400	1	1.483437	0.3983	264.76	0.0000
dv	400	2	2.89239	0.2207	113.26	0.0000

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
med						
iv	-.6360555	.0390902	-16.27	0.000	-.7126708	-.5594402
_cons	12.30566	.3954467	31.12	0.000	11.5306	13.08072
dv						
med	.8976328	.0974895	9.21	0.000	.7065569	1.088709
iv	.1641715	.0982558	1.67	0.095	-.0284064	.3567493
_cons	3.390934	1.426084	2.38	0.017	.5958621	6.186007

Correlation matrix of residuals:

	med	dv
med	1.0000	
dv	-0.0000	1.0000

Breusch-Pagan test of independence: chi2(1) = 0.000, Pr = 1.0000

The estimated correlation matrix among the error terms indicates the absence of significant correlation, which makes the SUR estimation unnecessary and the OLS estimate appropriate.

If, however, $\rho(\varepsilon_{1i}, \varepsilon_{2i}) \neq 0$, then Eq. (2) or Eq. (3) cannot be estimated through OLS because the coefficients are biased due to the correlation of the variable *med* with the error term. Taking equation (2) in its mean-centered form, the bias is determined by :

$$\hat{\beta}_{ols} = (\mathbf{med}'\mathbf{med})^{-1} \mathbf{med}'\mathbf{dv} \quad (5)$$

and

$$E[\hat{\beta}_{ols}] = E[(\mathbf{med}'\mathbf{med})^{-1} \mathbf{med}'\mathbf{dv}] = (\mathbf{med}'\mathbf{med})^{-1} E[\mathbf{med}'\mathbf{dv}] = (\mathbf{med}'\mathbf{med})^{-1} E[\mathbf{med}'\mathbf{dv}] \quad (6)$$

However

$$E[\mathbf{med}'\mathbf{dv}] = E[(\mathbf{iv}\boldsymbol{\alpha} + \boldsymbol{\varepsilon}_1)' (\mathbf{med}\boldsymbol{\beta} + \boldsymbol{\varepsilon}_2)] = (\mathbf{iv}\boldsymbol{\alpha})' (\mathbf{med}\boldsymbol{\beta}) E[\boldsymbol{\varepsilon}_1'\boldsymbol{\varepsilon}_2] \neq 0 \quad (7)$$

More specifically :

$$E[\boldsymbol{\varepsilon}_1'\boldsymbol{\varepsilon}_2] = \rho \quad (8)$$

This means that the OLS estimates as well as the SUR estimates are biased.

The correlation between the mediator and the error term biases not only the coefficient of the mediator, as shown above, but it biases all the model coefficients. This can be seen from the generalization of Eq. (5) to the vector of coefficients for the raw data (not mean centered).

Let

$$\mathbf{X}_{Tx3} = \begin{bmatrix} \boldsymbol{\iota} & med & iv \\ Tx1 & Tx1 & Tx1 \end{bmatrix} \quad (9)$$

where $\boldsymbol{\iota}$ is a column vector of ones.

Redefining \mathbf{dv} as \mathbf{y} , Eq. (3) in matrix notation becomes

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}_2 \quad (10)$$

The OLS estimator is

$$\hat{\boldsymbol{\beta}}_{ols} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{y} \quad (11)$$

Therefore,

$$E[\hat{\boldsymbol{\beta}}_{ols}] = (\mathbf{X}'\mathbf{X})^{-1} E[\mathbf{X}'\mathbf{y}] = (\mathbf{X}'\mathbf{X})^{-1} E[\mathbf{X}'(\mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}_2)] \quad (12)$$

$$E[\hat{\boldsymbol{\beta}}_{ols}] = \boldsymbol{\beta} + (\mathbf{X}'\mathbf{X})^{-1} E[\mathbf{X}'\boldsymbol{\varepsilon}_2] \quad (13)$$

Even if the error terms are uncorrelated to the unit vector and to \mathbf{iv} , the bias due to the correlation between *med* and the error term is transferred to the other parameters through the full covariance matrix $(\mathbf{X}'\mathbf{X})^{-1}$, the second part of Eq. (13).

$$\begin{aligned} (\mathbf{X}'\mathbf{X})^{-1} E[\mathbf{X}'\boldsymbol{\varepsilon}_2] &= (\mathbf{X}'\mathbf{X})^{-1} E \left[\begin{pmatrix} \boldsymbol{\iota}' \\ \mathbf{med}' \\ \mathbf{iv}' \end{pmatrix} \boldsymbol{\varepsilon}_2 \right] \\ &= (\mathbf{X}'\mathbf{X})^{-1} E \left[\begin{pmatrix} \boldsymbol{\iota}'\boldsymbol{\varepsilon}_2 \\ \mathbf{med}'\boldsymbol{\varepsilon}_2 \\ \mathbf{iv}'\boldsymbol{\varepsilon}_2 \end{pmatrix} \right] = (\mathbf{X}'\mathbf{X})^{-1} \begin{pmatrix} 0 \\ E[\mathbf{med}'\boldsymbol{\varepsilon}_2] \\ 0 \end{pmatrix} \neq 0 \end{aligned} \quad (14)$$

It should be noted that the $\mathbf{X}'\mathbf{X}$ matrix is not diagonal because, by Eq. (1) the mediator is correlated to the independent variable.

This bias can be seen by performing on OLS estimation on data generated as above but this time with a correlation of error terms across equations of 0.8. Table 3 shows that the parameters for the second equation are far from the parameters in the data generating function, in particular for the mediating variable (2.46 vs. the true parameter value of 0.76).

Table 3 – OLS estimation of correlated error model

```
. use "/Users/gatignon/Documents/WORK_STATA/MediationData/mediationdata_singleiv_Rho_08.dta",
clear
```

```
. regress med iv
```

Source	SS	df	MS	Number of obs = 400		
Model	727.109382	1	727.109382	F(1, 398)	=	361.38
Residual	800.793577	398	2.01204416	Prob > F	=	0.0000
				R-squared	=	0.4759
				Adj R-squared	=	0.4746
				Root MSE	=	1.4185

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
med						
iv	-.6632829	.0348913	-19.01	0.000	-.7318773	-.5946885
_cons	12.52571	.3650182	34.32	0.000	11.80811	13.24332


```
. regress dv med iv
```

Source	SS	df	MS	Number of obs = 400		
Model	5278.07946	2	2639.03973	F(2, 397)	=	682.87
Residual	1534.25683	397	3.86462676	Prob > F	=	0.0000
				R-squared	=	0.7748
				Adj R-squared	=	0.7736
				Root MSE	=	1.9659

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
dv						
med	2.455388	.0694694	35.34	0.000	2.318814	2.591962
iv	1.106728	.0667944	16.57	0.000	.9754126	1.238043
_cons	-15.27156	1.006521	-15.17	0.000	-17.25034	-13.29278

The SUR estimation does not even pick up the correlated errors as shown in Table 4.

Table 4 – SUR Estimation and correlation matrix of error terms

```

. sureg (med iv) (dv med iv), corr

Seemingly unrelated regression
-----
Equation      Obs  Parms      RMSE      "R-sq"      chi2      P
-----
med           400    1    1.414915    0.4759     363.19    0.0000
dv           400    2    1.95848    0.7748    1376.06    0.0000
-----

          |      Coef.  Std. Err.      z    P>|z|      [95% Conf. Interval]
-----+-----
med      |
      iv  |  -.6632829   .034804   -19.06  0.000   -1.7314975   -1.5950683
      _cons |  12.52571   .3641045   34.40  0.000   11.81208    13.23935
-----+-----
dv      |
      med  |   2.455388   .0692084   35.48  0.000    2.319742    2.591034
      iv   |   1.106728   .0665435   16.63  0.000    .9763048    1.23715
      _cons |  -15.27156   1.002739  -15.23  0.000   -17.23689   -13.30623
-----+-----

Correlation matrix of residuals:

      med      dv
med   1.0000
dv  -0.0000   1.0000

Breusch-Pagan test of independence: chi2(1) =      0.000, Pr = 1.0000

```

This is due to the fact that the correlation is picked up by the *iv* variable (with an estimated coefficient of 1.106 vs. the true parameter value of 0.1). Indeed, running a model that does not include *iv* in the second equation provides a correlation of the error terms of 0.44 (much closer to the true value of 0.8). These results are shown in Table 5.

Table 5 – SUR Estimation without *iv* variable and correlation matrix of error terms

```

. sureg (med iv) (dv med), corr

Seemingly unrelated regression
-----
Equation      Obs  Parms      RMSE      "R-sq"      chi2      P
-----
med           400    1    1.424719    0.4686     516.72    0.0000
dv           400    1    2.647516    0.5884     442.82    0.0000
-----

          |      Coef.  Std. Err.      z    P>|z|      [95% Conf. Interval]
-----+-----
med      |
      iv  |  -.7453697   .0327903  -22.73  0.000   -1.8096375   -1.6811019
      _cons |  13.3681    .3438565   38.88  0.000   12.69416    14.04205
-----+-----
dv      |
      med  |   1.291937   .0613941   21.04  0.000    1.171607    1.412267
      _cons |   2.739641   .3734964    7.34  0.000    2.007602    3.47168
-----+-----

Correlation matrix of residuals:

      med      dv
med   1.0000
dv   0.4411   1.0000

Breusch-Pagan test of independence: chi2(1) =     77.821, Pr = 0.0000

```

One can note that the bias due to the omitted variable *iv* is small in this case because the true parameter is only 0.1.

As a result of this analysis, we conclude that the mediator variable in Eq. (2) needs to be instrumented with an instrument that does not contain the correlated error term but that is highly correlated with this mediator.

One obvious candidate for instrument is the exogenous variable iv itself. Then Eq. (2) can be estimated by OLS or more efficiently by SUR :

$$\begin{cases} med_i = \alpha_0 + \alpha_1 iv_i + \varepsilon_{1i} \\ dv_i = \beta_0 + \beta_1 \hat{med}_i + \varepsilon_{2i} \end{cases} \quad (15)$$

This analysis is performed and the results are shown in Table 6.

Table 6 – Instrumental estimation

```

. regress med iv

      Source |           SS          df           MS          Number of obs =      400
-----+-----+-----+-----+-----+-----+-----+-----
      Model |    727.109382         1    727.109382          F( 1, 398) =    361.38
      Residual |    800.793577       398    2.01204416          Prob > F      =    0.0000
-----+-----+-----+-----+-----+-----+-----
      Total |   1527.90296       399    3.82933072          R-squared       =    0.4759
                                          Adj R-squared   =    0.4746
                                          Root MSE       =    1.4185

-----+-----+-----+-----+-----+-----+-----
      med |           Coef.      Std. Err.      t    P>|t|     [95% Conf. Interval]
-----+-----+-----+-----+-----+-----+-----
      iv |   -0.6632829      0.0348913   -19.01  0.000   -0.7318773   -0.5946885
      _cons |    12.52571      0.3650182    34.32  0.000    11.80811    13.24332
-----+-----+-----+-----+-----+-----+-----

. predict double medhat
(option xb assumed; fitted values)

. sureg (med iv) (dv medhat), corr

Seemingly unrelated regression
-----+-----+-----+-----+-----+-----+-----
Equation      Obs   Parms      RMSE      "R-sq"      chi2      P
-----+-----+-----+-----+-----+-----+-----
med           400     1     1.414915    0.4759     363.19    0.0000
dv           400     1     3.988165    0.0661     28.30     0.0000
-----+-----+-----+-----+-----+-----+-----

-----+-----+-----+-----+-----+-----+-----
      |           Coef.      Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----+-----+-----+-----+-----+-----
med   |
      iv |   -0.6632829      0.034804   -19.06  0.000   -0.7314975   -0.5950683
      _cons |    12.52571      0.3641045    34.40  0.000    11.81208    13.23935
-----+-----+-----+-----+-----+-----+-----
dv   |
      medhat |   0.7868274      0.1479017     5.32  0.000     0.4969453     1.076709
      _cons |   5.62835      0.8690341     6.48  0.000     3.925075     7.331626
-----+-----+-----+-----+-----+-----+-----

Correlation matrix of residuals:

      med      dv
med   1.0000
dv   0.8711  1.0000

Breusch-Pagan test of independence: chi2(1) = 303.539, Pr = 0.0000

. mat list e(Sigma)

symmetric e(Sigma)[2,2]
      med      dv
med   2.0019839
dv   4.915647  15.905462

```

The parameter estimates are extremely close to their true values, including in the second equation. The covariance of the error terms of 4.916 corresponds to a correlation of 0.87, which is significant and close to the true value of 0.8.

A problem arises, however, when attempting to estimate the model with the direct effect, i.e., Eq. (3). The reason is the perfect collinearity between iv and $m\hat{e}d$. Therefore, omitting iv in Eq. (3) leads also to a bias if $\beta_1 \neq 0$.

The analysis above highlights the need to incorporate covariates on the mediator and dependent equations. These covariates then act as exogenous variables. We discuss this case next.

Case of several exogenous (iv) variables

The case of several exogenous variables is similar. However, now it is possible to estimate a more complete model as long as one of the exogenous variables (i.e., one of the direct effects) is omitted from the last (dv) equation.

A new dataset was generated according to the equations in (16) with no correlation between the error terms:

$$\begin{aligned} med_i &= 12 - 0.6iv_{1i} + 0.3iv_{2i} + \varepsilon_{1i} \quad ; \quad \sigma_{\varepsilon_1} = 1.5 \\ dv_i &= 5 + 0.75med_i + 0.1iv_{1i} + \varepsilon_{2i} \quad ; \quad \sigma_{\varepsilon_2} = 3 \end{aligned} \tag{16}$$

The OLS and SUR estimations provide almost similar results, as the correlation is practically zero (i.e., -0.002) in Table 7.

Table 7 – Estimation of model with covariate (exogenous variable)

```
. use "/Users/gatignon/Documents/WORK_STATA/MediationData/mediationdata_MultipleIV_Rho_0.dta",
clear
```

```
. regress med iv1 iv2
```

Source	SS	df	MS	Number of obs =	400
Model	676.580788	2	338.290394	F(2, 397) =	135.21
Residual	993.273804	397	2.50194913	Prob > F =	0.0000
Total	1669.85459	399	4.18509923	R-squared =	0.4052
				Adj R-squared =	0.4022
				Root MSE =	1.5818

med	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
iv1	-.5840254	.0398751	-14.65	0.000	-.6624181 -.5056327
iv2	.2797617	.0400665	6.98	0.000	.2009928 .3585307
_cons	12.00387	.5830706	20.59	0.000	10.85758 13.15016

```
. regress dv med iv1
```

Source	SS	df	MS	Number of obs =	400
Model	903.523301	2	451.761651	F(2, 397) =	41.84
Residual	4287.03229	397	10.79857	Prob > F =	0.0000
Total	5190.5556	399	13.0089113	R-squared =	0.1741
				Adj R-squared =	0.1699
				Root MSE =	3.2861

dv	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
med	.75889	.0984002	7.71	0.000	.5654394 .9523406
iv1	.043062	.1013104	0.43	0.671	-.1561099 .2422338
_cons	5.563675	1.695621	3.28	0.001	2.230156 8.897194

```
. sureg (med iv1 iv2) (dv med iv1), corr
```

Seemingly unrelated regression

Equation	Obs	Parms	RMSE	"R-sq"	chi2	P
med	400	2	1.575812	0.4052	272.47	0.0000
dv	400	2	3.273778	0.1741	84.90	0.0000

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
med					
iv1	-.5840251	.0397252	-14.70	0.000	-.6618852 -.5061651
iv2	.2797703	.0399159	7.01	0.000	.2015366 .358004
_cons	12.00378	.5808794	20.66	0.000	10.86528 13.14229
dv					
med	.7626875	.0980303	7.78	0.000	.5705515 .9548234
iv1	.0453152	.1009297	0.45	0.653	-.1525034 .2431337
_cons	5.507104	1.689248	3.26	0.001	2.196238 8.81797

Correlation matrix of residuals:

	med	dv
med	1.0000	
dv	-0.0021	1.0000

Breusch-Pagan test of independence: chi2(1) = 0.002, Pr = 0.9673

```
. mat list e(Sigma)
```

```
symmetric e(Sigma)[2,2]
      med      dv
med 2.4831845
```


dv **-.01058788** **10.717581**

As in the simple case of a single *iv*, the parameters are reproduced relatively accurately.

Another dataset was generated with the same parameter values as in Eq. (16) but with a correlation of error terms across equation for 0.8.

The OLS or SUR estimations clearly show biased parameter values for the second equation (Table 8). For example the coefficient of the mediator is 2.14 (OLS) and 1.95 (SUR) instead of the true value of 0.76.

Table 8 – OLS and SUR estimations

```
. use
"/Users/gatignon/Documents/WORK_STATA/MediationData/mediationdata_MultipleIV_Rho_08.dta",
clear
```

```
. regress med iv1 iv2
```

Source	SS	df	MS	Number of obs =	400
Model	646.207781	2	323.10389	F(2, 397) =	142.33
Residual	901.250972	397	2.27015358	Prob > F =	0.0000
				R-squared =	0.4176
				Adj R-squared =	0.4147
Total	1547.45875	399	3.87834274	Root MSE =	1.5067

med	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
iv1	-.5751035	.039274	-14.64	0.000	-.6523146 -.4978925
iv2	.3489572	.0384289	9.08	0.000	.2734077 .4245067
_cons	11.16968	.5361259	20.83	0.000	10.11568 12.22368

```
. regress dv med iv1
```

Source	SS	df	MS	Number of obs =	400
Model	5067.22472	2	2533.61236	F(2, 397) =	663.03
Residual	1517.04826	397	3.82128024	Prob > F =	0.0000
				R-squared =	0.7696
				Adj R-squared =	0.7684
Total	6584.27298	399	16.5019373	Root MSE =	1.9548

dv	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
med	2.14051	.0592518	36.13	0.000	2.024023 2.256997
iv1	.9607916	.0606844	15.83	0.000	.8414886 1.080095
_cons	-15.93324	.9985147	-15.96	0.000	-17.89627 -13.9702

```
. sureg (med iv1 iv2) (dv med iv1), corr
```

Seemingly unrelated regression

Equation	Obs	Parms	RMSE	"R-sq"	chi2	P
med	400	2	1.503605	0.4156	312.38	0.0000
dv	400	2	1.971756	0.7638	1146.28	0.0000

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
med					
iv1	-.5773227	.0391252	-14.76	0.000	-.6540067 -.5006387
iv2	.3937164	.0377711	10.42	0.000	.3196865 .4677464
_cons	10.74818	.5308602	20.25	0.000	9.707716 11.78865
dv					
med	1.953465	.0582376	33.54	0.000	1.839321 2.067608
iv1	.8564574	.0602171	14.22	0.000	.738434 .9744807
_cons	-13.22935	.9849646	-13.43	0.000	-15.15984 -11.29885

Correlation matrix of residuals:

	med	dv
med	1.0000	
dv	0.1789	1.0000

Breusch-Pagan test of independence: chi2(1) = 12.799, Pr = 0.0003

```
. mat list e(Sigma)
```

```
symmetric e(Sigma)[2,2]
      med      dv
med 2.2531274
```

dv .52290004 3.7926206

When applying the instrumental approach, however, the parameters are recovered much more accurately. In Table 9, the mediator variable is predicted first using the two exogenous variables as predictors and the predicted value of the mediator is used in the second equation as an instrument for the mediator variable. SUR is used to account for the correlated errors across equations.

Table 9 – Instrumental estimation

```
. regress med iv1 iv2
```

Source	SS	df	MS	Number of obs =	400
Model	646.207781	2	323.10389	F(2, 397) =	142.33
Residual	901.250972	397	2.27015358	Prob > F =	0.0000
				R-squared =	0.4176
				Adj R-squared =	0.4147
Total	1547.45875	399	3.87834274	Root MSE =	1.5067

med	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
iv1	-.5751035	.039274	-14.64	0.000	-.6523146 -.4978925
iv2	.3489572	.0384289	9.08	0.000	.2734077 .4245067
_cons	11.16968	.5361259	20.83	0.000	10.11568 12.22368

```
. predict medhat
(option xb assumed; fitted values)
```

```
. sureg (med iv1 iv2) (dv medhat iv1), corr
```

Seemingly unrelated regression

Equation	Obs	Parms	RMSE	"R-sq"	chi2	P
med	400	2	1.501042	0.4176	286.80	0.0000
dv	400	2	3.971176	0.0419	17.51	0.0002

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
med					
iv1	-.5751035	.0391264	-14.70	0.000	-.65179 -.4984171
iv2	.3489573	.0382845	9.11	0.000	.2739211 .4239935
_cons	11.16968	.5341116	20.91	0.000	10.12284 12.21652
dv					
medhat	1.023148	.2902531	3.53	0.000	.4542622 1.592033
iv1	.3375245	.1921007	1.76	0.079	-.0389859 .7140348
_cons	.2191332	4.323431	0.05	0.960	-8.254636 8.692903

Correlation matrix of residuals:

	med	dv
med	1.0000	
dv	0.8968	1.0000

Breusch-Pagan test of independence: chi2(1) = 321.701, Pr = 0.0000

```
. mat list e(Sigma)
```

```
symmetric e(Sigma)[2,2]
```

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```

      med      dv
med  2.2531274
dv   5.3457417  15.770238

```

These steps are performed automatically with a three stage least squares estimation and the results are identical as shown in Table 10.

Table 10 – Three-stage least squares estimates

```

. reg3 (med iv1 iv2) (dv med iv1)

Three-stage least-squares regression
-----
Equation      Obs   Parns      RMSE      "R-sq"      chi2      P
-----
med           400     2    1.501042    0.4176    286.80    0.0000
dv           400     2    2.681402    0.5632     38.41    0.0000
-----

      Coef.   Std. Err.      z    P>|z|      [95% Conf. Interval]
-----+-----
med
   iv1    -0.5751035   0.0391264   -14.70   0.000   -0.65179   -0.4984171
   iv2     0.3489572   0.0382845    9.11   0.000    0.273921   0.4239934
   _cons   11.16968     0.5341116   20.91   0.000   10.12284   12.21652
-----
dv
   med     1.023147   0.1959836    5.22   0.000    0.6390266   1.407268
   iv1    -0.3375242   0.1297095    2.60   0.009   -0.0832983   0.5917501
   _cons    0.21914    2.919251    0.08   0.940   -5.502486   5.940766
-----

Endogenous variables:  med dv
Exogenous variables:  iv1 iv2
-----

. mat list e(Sigma)

symmetric e(Sigma)[2,2]
      med      dv
med  2.2531274
dv   3.0404604  7.1899174

```

Note that the correlation estimate is slightly different with values of 0.89 and 0.76 (as obtained from the estimated covariance matrix Sigma) respectively for the estimates shown in Table 9 and in Table 10.

A final remark can be useful concerning the example found in <http://www.ats.ucla.edu/stat/stata/faq/mulmediation.htm>. The SUR estimates are reported as shown in Table 11.

Table 11 – SUR estimates of education data

```

. use http://www.ats.ucla.edu/stat/data/hsb2, clear
(highschool and beyond (200 cases))

. sureg (read math)(write math)(science read write math), corr

```

Seemingly unrelated regression

Equation	Obs	Parms	RMSE	"R-sq"	chi2	P
read	200	1	7.662848	0.4386	156.26	0.0000
write	200	1	7.437294	0.3812	123.23	0.0000
science	200	3	6.983853	0.4999	199.96	0.0000

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
read					
math	.724807	.0579824	12.50	0.000	.6111636 .8384504
_cons	14.07254	3.100201	4.54	0.000	7.996255 20.14882
write					
math	.6247082	.0562757	11.10	0.000	.5144099 .7350065
_cons	19.88724	3.008947	6.61	0.000	13.98981 25.78467
science					
read	.3015317	.0679912	4.43	0.000	.1682715 .434792
write	.2065257	.0700532	2.95	0.003	.0692239 .3438274
math	.3190094	.0759047	4.20	0.000	.170239 .4677798
_cons	8.407353	3.160709	2.66	0.008	2.212476 14.60223

Correlation matrix of residuals:

	read	write	science
read	1.0000		
write	0.3187	1.0000	
science	-0.0000	0.0000	1.0000

Breusch-Pagan test of independence: chi2(3) = 20.318, Pr = 0.0001

```

. mat list e(Sigma)

```

symmetric e(Sigma)[3,3]

	read	write	science
read	58.719247		
write	18.164815	55.313342	
science	-1.839e-13	3.930e-13	48.77421

However, it is not possible to apply the instrumental methods because the third equation is not identified. In this particularly example, the structure of the error terms covariance is block diagonal where the error term in the third equation is uncorrelated with the error terms in the other two equations, as shown in the correlation matrix of residuals.

Consequently, the SUR estimation shown in Table 11 is not biased. This would not be generally the case, however, with correlated error terms.

In summary, this shows that identification is necessary in recursive systems when the equation errors are correlated and that instrumental variables can recover the true parameter estimates. However, this requires that covariates are included in the analysis. When the equation errors are not correlated, separate OLS estimation of each equation is possible.

SUR vs. ISUR

The iterative SUR estimation re-estimates the covariance matrix of error terms until the parameter estimates converge. These estimates are the maximum likelihood estimates.

We use the same model as in Table 9, where the SUR estimates using the instrumental variable approach are provided. The ISUR estimates are shown on the top part of Table 12. The log-likelihood is also provided upon request with the command “**display e(ll)**”, or using the command “**estat ic**” that gives, in addition to the log-likelihood, Akaike’s AIC criterion as well as the BIC criterion. These measures that be used to compare non-nested models since they correct the log-likelihood for the number of parameters estimated. The coefficient estimates are exactly the same as the SUR estimates. This is often the case in simple models that converge rapidly. However, the ISUR parameters (maximum likelihood) are asymptotically more efficient.

Table 12 – ISUR estimates and estimates using SEM

```
. *ISUR
. sureg (med iv1 iv2) (dv medhat iv1) ///
> , isure corr nolog
```

Seemingly unrelated regression, iterated

Equation	Obs	Parms	RMSE	"R-sq"	chi2	P
med	400	2	1.501042	0.4176	286.80	0.0000
dv	400	2	3.971176	0.0419	17.51	0.0002

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
med					
iv1	-.5751035	.0391264	-14.70	0.000	-.65179 -.4984171
iv2	.3489573	.0382845	9.11	0.000	.2739211 .4239935
_cons	11.16968	.5341116	20.91	0.000	10.12284 12.21652
dv					
medhat	1.023148	.2902531	3.53	0.000	.4542622 1.592033
iv1	.3375245	.1921007	1.76	0.079	-.0389859 .7140348
_cons	.2191332	4.323431	0.05	0.960	-8.254636 8.692903

Correlation matrix of residuals:

	med	dv
med	1.0000	
dv	0.8968	1.0000

Breusch-Pagan test of independence: chi2(1) = 321.701, Pr = 0.0000

```
. display e(ll)
-1523.0545
```

```
. estat ic
```

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	400	.	-1523.055	6	3058.109	3082.058

Note: N=Obs used in calculating BIC; see [R] BIC note.

```
. *SEM
. sem (med <- iv1 iv2) (dv <- medhat iv1) ///
> , cov(e.med*e.dv )
```

Endogenous variables

Observed: med dv

Exogenous variables

Observed: iv1 iv2 medhat

Fitting target model:

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```

Iteration 0:   log likelihood =   2327.341
Iteration 1:   log likelihood =   2327.4698   (backed up)

Structural equation model                               Number of obs   =       400
Estimation method = ml
Log likelihood   =   2327.4698

-----+-----
              |              OIM
              |      Coef.    Std. Err.      z    P>|z|      [95% Conf. Interval]
-----+-----+-----
Structural
med <-
   iv1       |  -.5751035   .0402632   -14.28   0.000   -.6540179   -.4961892
   iv2       |   .3489572   .0376988    9.26    0.000    .275069    .4228454
   _cons     |  11.16968   .5238414   21.32   0.000   10.14297   12.19639
-----+-----
dv <-
   iv1       |   .3375242   .1864021    1.81    0.070   -.0278172   .7028657
   medhat    |   1.023147   .2850764    3.59    0.000    .4644079   1.581887
   _cons     |   .2191389   4.202358    0.05    0.958   -8.017331   8.455608
-----+-----
var(e.med)   |   2.253127   .1593219                                1.961535   2.588067
var(e.dv)   |   15.77024   1.115133                                13.72931   18.11456
-----+-----
cov(e.med,e.dv) |   5.345742   .4003458   13.35   0.000   4.561078   6.130405
-----+-----
LR test of model vs. saturated: chi2(2)   =       1.93, Prob > chi2 = 0.3800

```

SUR vs. SEM

The question we address in this section concerns the comparison of results obtained from SUR versus those obtained via SEM.

First, we will use the same model specification as above again, i.e., using the instrumental variable approach. The parameter estimated with SEM are shown at the bottom of Table 12. Again, the estimated coefficients are identical up to the sixth digit after the decimal point. Note that we requested the covariance between the error terms of the two equations to be estimated (this issue is discussed next). This covariance matrix is identical to the one obtained with the SUR or ISUR estimation methods. However, the log-likelihood that results from this estimation is different from the one obtained with the ISUR estimation (2327.4698 vs. -1523.0545). These two log-likelihood are not comparable because the likelihood functions is different with the two models. In the case of SEM, the likelihood function is relative to the likelihood under the assumption of the saturated model, which explains the positive value provided in the output for the SEM estimation. (See Gatignon, 2014, pages 79-82).

Not knowing a priori if the error terms of each equation are correlated or not, we have to assume that the model is identified, which is the only way to get proper parameter estimates of the covariances among the error terms.

We will now use a model specification slightly more complex with one independent variable (iv) that is the focus of the study, two mediating variables (med1 and med2) leading to two mediator equations, and two control variables (cv1 and cv2). Therefore, the

model has three exogenous variables (iv, cv1 and cv2), and three endogenous variables (med1, med2 and dv). The data used for illustrating the difference between SUR with Instruments (or 3SLS) versus SEM contains a significant correlation (0.8) between the error terms of the second mediator equation and the dv equation). Note that SEM is used here only to reflect structural relationships, as no measurement model is specified and each variable is assumed to be observed without error, just as in the regression examples above; the estimation of an SEM model with both measurement and structural parameters is covered in Chapter 11 of Gatignon (2014).

First we apply the SUR estimation with Instrumental variables in the dv equation for which the code is shown in shown in Table 13:

Table 13 – SUR estimates with instrumented mediator variables-code

```
regress med1 iv cv1 cv2
predict med1hat
regress med2 iv cv1 cv2
predict med2hat
sureg (med1 iv cv1 cv2) (med2 iv cv1 cv2) (dv iv med1hat med2hat), corr
mat list e(Sigma)
```

The results are shown in Table 14:

Table 14 – SUR estimates with instrumented mediator variables-results

```
. regress med1 iv cv1 cv2
```

Source	SS	df	MS	Number of obs	=	400
Model	760.269157	3	253.423052	F(3, 396)	=	119.18
Residual	842.072787	396	2.12644643	Prob > F	=	0.0000
				R-squared	=	0.4745
				Adj R-squared	=	0.4705
Total	1602.34194	399	4.0158946	Root MSE	=	1.4582

med1	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
iv	.4353785	.0360196	12.09	0.000	.3645649 .5061921
cv1	-.5158117	.0387398	-13.31	0.000	-.591973 -.4396504
cv2	-.0141616	.0335059	-0.42	0.673	-.0800332 .05171
_cons	10.87417	.6392709	17.01	0.000	9.61738 12.13096


```
. predict med1hat
(option xb assumed; fitted values)
```

```
. regress med2 iv cv1 cv2
```

Source	SS	df	MS	Number of obs	=	400
Model	570.653089	3	190.217696	F(3, 396)	=	49.32
Residual	1527.25292	396	3.85669928	Prob > F	=	0.0000
				R-squared	=	0.2720
				Adj R-squared	=	0.2665
Total	2097.90601	399	5.25790979	Root MSE	=	1.9638

med2	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
iv	.5143321	.0485087	10.60	0.000	.4189653 .6096989
cv1	.0337605	.052172	0.65	0.518	-.0688083 .1363293
cv2	.2741464	.0451234	6.08	0.000	.1854351 .3628576
_cons	7.657081	.8609255	8.89	0.000	5.964525 9.349637


```
. predict med2hat
(option xb assumed; fitted values)
```

```
. sureg (med1 iv cv1 cv2) (med2 iv cv1 cv2) (dv iv med1hat med2hat), corr
```

Seemingly unrelated regression

Equation	Obs	Parms	RMSE	"R-sq"	chi2	P
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```

-----
med1      400      3      1.450925      0.4745      361.14      0.0000
med2      400      3      1.954004      0.2720      149.46      0.0000
dv        400      3      2.388436      0.1846      90.53      0.0000
-----

```

```

-----
          |      Coef.      Std. Err.      z      P>|z|      [95% Conf. Interval]
-----+-----
med1      |
   iv     |      .4353785     .0358391     12.15     0.000     .3651352     .5056218
   cv1    |     -.5158117     .0385456    -13.38     0.000    -.5913596    -.4402637
   cv2    |     -.0141617     .0333379     -0.42     0.671    -.0795028     .0511794
   _cons  |     10.87417      .6360665     17.10     0.000     9.627501     12.12084
-----+-----
med2      |
   iv     |      .5143321     .0482656     10.66     0.000     .4197334     .6089309
   cv1    |      .0337605     .0519105      0.65     0.515    -.0679822     .1355032
   cv2    |      .2741463     .0448972      6.11     0.000     .1861495     .3621431
   _cons  |      7.657081     .8566101      8.94     0.000     5.978157     9.336006
-----+-----
dv        |
   iv     |      .2742002     .1412007      1.94     0.052    -.0025482     .5509485
 med1hat  |      .8005514     .1250243      6.40     0.000     .5555082     1.045595
 med2hat  |     -.7266239     .2013998     -3.61     0.000    -1.12136    -.3318876
   _cons  |      9.386481     2.471676      3.80     0.000     4.542085     14.23088
-----

```

Correlation matrix of residuals:

```

          med1      med2      dv
med1     1.0000
med2    -0.0320      1.0000
dv       0.4595      0.5215      1.0000

```

Breusch-Pagan test of independence: $\chi^2(3) = 193.637$, Pr = 0.0000

. mat list e(Sigma)

```

symmetric e(Sigma)[3,3]
          med1      med2      dv
med1     2.105182
med2    -.09064448      3.8181323
dv       1.5923296      2.4337537      5.7046267

```

As shown earlier in this note, the estimated coefficients are the same as when using 3SLS, as they are shown in Table 15:

Table 15 – 3SLS estimates

```

. reg3 (med1 iv cv1 cv2) (med2 iv cv1 cv2) (dv iv med1 med2)

Three-stage least-squares regression
-----
Equation          Obs    Parns      RMSE      "R-sq"      chi2      P
-----
med1              400      3    1.450925    0.4745     361.14    0.0000
med2              400      3    1.954004    0.2720     149.46    0.0000
dv                400      3    3.187871   -0.4527     50.82    0.0000
-----

          |      Coef.   Std. Err.   z    P>|z|   [95% Conf. Interval]
-----+-----
med1      |
   iv     |   .4353785   .0358391   12.15  0.000   .3651352   .5056218
   cv1    |  -.5158117   .0385456  -13.38  0.000  -.5913596  -.4402637
   cv2    |  -.0141616   .0333379   -0.42  0.671  -.0795027   .0511795
   _cons  |   10.87417   .6360665   17.10  0.000   9.627501  12.12084
-----+-----
med2      |
   iv     |   .5143321   .0482656   10.66  0.000   .4197334   .6089309
   cv1    |   .0337605   .0519105    0.65  0.515  -.0679822   .1355032
   cv2    |   .2741464   .0448972    6.11  0.000   .1861495   .3621432
   _cons  |   7.657081   .8566101    8.94  0.000   5.978156  9.336006
-----+-----
dv        |
   iv     |   .2741999   .1884621    1.45  0.146  -.0951791   .6435789
   med1   |   .8005515   .1668713    4.80  0.000   .4734897   1.127613
   med2   |  -.7266235   .2688104   -2.70  0.007  -1.253482  -.1997647
   _cons  |   9.386476   3.298973    2.85  0.004   2.920608  15.85234
-----+-----

Endogenous variables:  med1 med2 dv
Exogenous variables:  iv cv1 cv2
-----

. mat list e(Sigma)

symmetric e(Sigma)[3,3]
      med1      med2      dv
med1   2.105182
med2  -.09064448   3.8181323
dv   -.15884142   5.2806639  10.162523

```

When we use the SEM procedure, we obtain the estimates shown in Table 16:

Table 16 – SEM model and estimated parameters

```

. sem (med1 <- iv cv1 cv2) (med2 <- iv cv1 cv2) (dv <- iv med1 med2)

Endogenous variables

Observed:  med1 med2 dv

Exogenous variables

Observed:  iv cv1 cv2

Fitting target model:

Iteration 0:  log likelihood = -4921.2392
Iteration 1:  log likelihood = -4921.2392

Structural equation model          Number of obs    =      400
Estimation method = ml
Log likelihood = -4921.2392

-----
          |      Coef.      OIM      Std. Err.   z    P>|z|   [95% Conf. Interval]
-----+-----
Structural
med1 <-  |
   iv     |   .4353785   .0358391   12.15  0.000   .3651352   .5056218
   cv1    |  -.5158117   .0385456  -13.38  0.000  -.5913596  -.4402637

```

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```

      cv2 | -.0141616 .0333379 -0.42 0.671 -.0795027 .0511795
      _cons | 10.87417 .6360665 17.10 0.000 9.627501 12.12084
-----+-----
med2 <-
      iv | .5143321 .0482656 10.66 0.000 .4197334 .6089309
      cv1 | .0337605 .0519105 0.65 0.515 -.0679822 .1355032
      cv2 | .2741464 .0448972 6.11 0.000 .1861495 .3621432
      _cons | 7.657081 .8566101 8.94 0.000 5.978156 9.336006
-----+-----
dv <-
      med1 | .8359019 .0535579 15.61 0.000 .7309302 .9408735
      med2 | .5369659 .045776 11.73 0.000 .4472465 .6266852
      iv | -.3826769 .0581491 -6.58 0.000 -.4966471 -.2687068
      _cons | -4.441153 .7549511 -5.88 0.000 -5.92083 -2.961476
-----+-----
var(e.med1) | 2.105182 .1488588 1.832739 2.418125
var(e.med2) | 3.818132 .2699827 3.324007 4.385711
var(e.dv) | 3.495547 .2471725 3.043169 4.015173
-----+-----
LR test of model vs. saturated: chi2(3) = 80.88, Prob > chi2 = 0.0000

```

```
. estat gof
```

Fit statistic	Value	Description
Likelihood ratio		
chi2_ms(3)	80.876	model vs. saturated
p > chi2	0.000	
chi2_bs(12)	742.732	baseline vs. saturated
p > chi2	0.000	

The parameters for the effects of the mediators and the direct effect if iv on the dv are clearly different from those obtained using SUR with the instrumented mediators, and so are the variance of the error term of the third equation.

What is not clear is what assumption is used for the covariances among these three error terms. We then run an SEM model where the covariances are constrained to 0 and another one where they are specifically estimated. Tables 17 shows the results where the covariances are constrained to 0:

Table 17 – SEM model with covariances of error terms constrained to 0

```

. sem (med1 <- iv cv1 cv2) (med2 <- iv cv1 cv2) (dv <- iv med1 med2) ///
> , cov(e.med1*e.med2@0) cov(e.med1*e.dv@0) cov(e.med2*e.dv@0)

Endogenous variables
Observed: med1 med2 dv

Exogenous variables
Observed: iv cv1 cv2

Fitting target model:

Iteration 0: log likelihood = -4921.2392
Iteration 1: log likelihood = -4921.2392

Structural equation model          Number of obs   =       400
Estimation method = ml
Log likelihood = -4921.2392

```

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	Coef.	OIM Std. Err.	z	P> z	[95% Conf. Interval]	
Structural						
med1 <-						
iv	.4353785	.0358391	12.15	0.000	.3651352	.5056218
cv1	-.5158117	.0385456	-13.38	0.000	-.5913596	-.4402637
cv2	-.0141616	.0333379	-0.42	0.671	-.0795027	.0511795
_cons	10.87417	.6360665	17.10	0.000	9.627501	12.12084
med2 <-						
iv	.5143321	.0482656	10.66	0.000	.4197334	.6089309
cv1	.0337605	.0519105	0.65	0.515	-.0679822	.1355032
cv2	.2741464	.0448972	6.11	0.000	.1861495	.3621432
_cons	7.657081	.8566101	8.94	0.000	5.978156	9.336006
dv <-						
med1	.8359019	.0535579	15.61	0.000	.7309302	.9408735
med2	.5369659	.045776	11.73	0.000	.4472465	.6266852
iv	-.3826769	.0581491	-6.58	0.000	-.4966471	-.2687068
_cons	-4.441153	.7549511	-5.88	0.000	-5.92083	-2.961476
var(e.med1)	2.105182	.1488588			1.832739	2.418125
var(e.med2)	3.818132	.2699827			3.324007	4.385711
var(e.dv)	3.495547	.2471725			3.043169	4.015173

LR test of model vs. saturated: $\chi^2(3) = 80.88$, Prob > $\chi^2 = 0.0000$

. estat gof

Fit statistic	Value	Description
Likelihood ratio		
chi2_ms(3)	80.876	model vs. saturated
p > chi2	0.000	
chi2_bs(12)	742.732	baseline vs. saturated
p > chi2	0.000	

These results are identical to those when nothing is specified about the covariances of the error terms, which shows that failure to specifically ask for their estimation assumes they are independent, which is not appropriate if this is not the case. Consequently the model should be specified with these covariances as parameters to be estimated. Table 18 shows the results where the covariances of the error terms are estimated:

Table 18 – SEM model with estimated covariances of error terms

```
. sem (med1 <- iv cv1 cv2) (med2 <- iv cv1 cv2) (dv <- iv med1 med2) ///
> , cov(e.med1*e.med2) cov(e.med1*e.dv) cov(e.med2*e.dv )

Endogenous variables

Observed:  med1 med2 dv

Exogenous variables

Observed:  iv cv1 cv2

Fitting target model:

Iteration 0:  log likelihood = -4880.8012
Iteration 1:  log likelihood = -4880.8012

Structural equation model                Number of obs    =        400
Estimation method    = ml
Log likelihood       = -4880.8012
```

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		OIM				
		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
Structural						
med1 <-						
	iv	.4353785	.0358391	12.15	0.000	.3651352 .5056218
	cv1	-.5158117	.0385456	-13.38	0.000	-.5913596 -.4402637
	cv2	-.0141616	.0333379	-0.42	0.671	-.0795027 .0511795
	_cons	10.87417	.6360665	17.10	0.000	9.627501 12.12084
med2 <-						
	iv	.5143321	.0482656	10.66	0.000	.4197334 .6089309
	cv1	.0337605	.0519105	0.65	0.515	-.0679822 .1355032
	cv2	.2741464	.0448972	6.11	0.000	.1861495 .3621432
	_cons	7.657081	.8566101	8.94	0.000	5.978156 9.336006
dv <-						
	med1	.8005515	.1668713	4.80	0.000	.4734897 1.127613
	med2	-.7266235	.2688104	-2.70	0.007	-1.253482 -.1997647
	iv	.2741999	.1884621	1.45	0.146	-.0951791 .6435789
	_cons	9.386476	3.298973	2.85	0.004	2.920608 15.85234
	var(e.med1)	2.105182	.1488588			1.832739 2.418125
	var(e.med2)	3.818132	.2699827			3.324007 4.385711
	var(e.dv)	10.16252	2.918504			5.788298 17.84236
	cov(e.med1,e.med2)	-.0906445	.1418281	-0.64	0.523	-.3686224 .1873334
	cov(e.med1,e.dv)	-.1588414	.4172006	-0.38	0.703	-.9765395 .6588567
	cov(e.med2,e.dv)	5.280664	1.101823	4.79	0.000	3.121131 7.440197

LR test of model vs. saturated: chi2(0) = 0.00, Prob > chi2 = .

. estat gof

Fit statistic	Value	Description
Likelihood ratio		
chi2_ms(0)	0.000	model vs. saturated
p > chi2	.	
chi2_bs(12)	742.732	baseline vs. saturated
p > chi2	0.000	

The coefficients obtained with this model specification are exactly the same as with the SUR using instrumented mediators and as in the 3SLS estimation. The variances and covariances are also exactly the same as those obtained in 3SLS.

In conclusion, SEM gives the same results but it is important to remember to free the covariances among the error terms to be estimated. One issue, however, that appears in these examples, concerns the measures of goodness of fit in the SEM approach. SEM estimation is an analysis of covariance structure where the data consists in the variances and covariances among the exogenous and endogenous variables. Consequently, in Structural Equation Models that do not include multiple item measures, the degrees of freedom are small at best. Our model where we estimate the covariance is in fact saturated and has no degrees of freedom at all. The significance of the models obtained from least squares estimations can be determined based on the usual degrees of freedom that are based on the number of observations, even if the R-squared are not interpretable for generalized least squares estimated models since they vary between minus infinity and 1 (Judge et al. 1985).

References

Gatignon, Hubert (2014), *Statistical Analysis of Management Data*, Springer Science+Business Media, LLC, New York, NY.

Judge, George G., William E. Griffiths, R. Carter Hill, Helmut Lütkepohl, and Tsoung-Chao Lee (1985), *The Theory and Practice of Econometrics*, 2nd Edition, New York, NY: John Wiley and Sons, Inc..

“Stata FAQ: How Can I Analyze Multiple Mediators in Stata?” (n.d.), <<http://www.ats.ucla.edu/stat/stata/faq/mulmediation.htm>> (Jul. 7, 2013).